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# **Prediction and Interpretation for Machine Learning Regression Methods**

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## **ABSTRACT**

The last 30 years has seen extraordinary development of new tools for the prediction of numerical and binary responses. Examples include the LASSO and elastic net for regularization in regression and variable selection, quantile regression for heteroscedastic data, and machine learning predictive method such as classification and regression trees (CART), multivariate adaptive regression splines (MARS), random forests, gradient boosting machines (GBM), and support vector machines (SVM). All these methods are implemented in SAS®, giving the user an amazing toolkit of predictive methods. In fact, the set of available methods is so rich it begs the question, "When should I use one or a subset of these methods instead of the other methods?" In this talk I hope to provide a partial answer to this question through the application of several of these methods in the analysis of several real datasets with numerical and binary response variables.

#### INTRODUCTION

Over the last 30 years there has been substantial development of regression methodology for regularization of the estimation in the multiple linear regression model and for carrying out non-linear regression of various kinds. Notable contributions in the area of regularization include the LASSO (Tibshirani 1996), the elastic net (Zou and Hastie 2005), and least angle regression (Effron et al. 2002) which is both a regularization method and a series of algorithms that can be used to efficiently compute LASSO and elastic net estimates of regression coefficients.

An early paper on non-linear regression via scatter plot smoothing and the alternating conditional expectations (ACE) algorithm is due to Breiman and Friedman (1985). Hastie and Tibshirani (1986) extend this approach to create generalized additive models (GAM). An alternative approach to non-linear regression using binary partitioning are regression trees (Breiman et al. 1984). Multivariate adaptive regression splines (MARS) (Friedman 1991) extended generalized linear and generalized additive models in the direction of modeling interactions, and considerable research of tree methods, notably ensembles of trees, resulted in the development of gradient boosting machines (GBM) (Friedman 2000) and random forests (Breiman 2001). A completely different approach, based on non-linear projections is support vector machines, the modern development of which is usually credited to Vapnik (1995) and Cortes and Vapnik (1995).

All of the methods listed above, and more, are implemented in SAS and other statistical packages giving statisticians a very large toolkit for analyzing and understanding data with a continuous (interval valued) response variable. In SAS using the LASSO or fitting a regression tree or random forests is no harder than fitting an ordinary multiple regression with some traditional variable selection. The LASSO has rapidly become a "standard" method for variable selection in regression, and all of these methods lend themselves to larger datasets, where there is a lot of information and statistical significance does not make sense.

In this paper I hope to illustrate the use of some of these methods for the analysis of real datasets.

#### **GETTING STARTED**

In the spirit of the "Getting Started" section of SAS procedure manual entries, we begin with a simple example that illustrates how tree methods can provide insight in situations where linear methods are less effective. The data concern credit card applications to a bank in Australia (Quinlan, 1987). The response variable is coded as "Yes" if the application was approved and "No" if it was not approved. There are 15 predictor variables denoted by A1—A15, some categorical and some numerical. For proprietary reasons the nature of the variables is not available. We note that variables A9 and A10 are code as 't' and 'f' which we take to mean 'true' and 'false.' A total of 666 observations had no missing values and of those 299 persons were approved for credit cards and 367 were not.

A first step in a traditional analysis might be to fit a logistic regression, perhaps with some form of variable selection. For this example I used backward elimination with a significance level to stay of  $\alpha = 0.05$ . The code is given below:

Eight variables were removed from the model. From the output for the ctable option we obtain the classification accuracy metrics for the fitted model.

Error! Reference source not found.. Classification accuracy for logistics regression on credit card approval data.

Classification Table									
	Correct Incorrect Percentages								
Prob Level	Event	Non- Event	Event	Non- Event	Correct	Sensitivity	Specificity	False POS	False NEG
0.500	273	319	57	28	87.4	90.7	84.8	17.3	8.1

The accuracy of the predictions is quite good with an overall percent correct of 87.4% (which means the overall error rate is 12.6% = 0.126), and both the sensitivity (percent of approval correctly predicted) and specificity (percent of non-approvals correctly predicted) quite high at 90.7% and 84.8%, respectively. The receiving operating characteristic (ROC) curve is a graphical representation of the quality of the fit of predictive model for a binary target (response). Figure 1 shows the ROC curves for all the steps in the variable elimination process overlaid. It is clear from this graph that the variables eliminated from the model were not contributing to the predictive accuracy and that the overall fit of the logistic regression model is rather good. The AUC value of 0.9463 for the model is high.

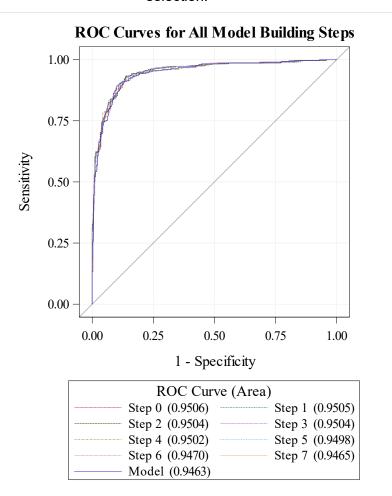


Table 2 contains the estimated coefficients for the variable remaining in the model. From this table it is relatively difficult to tell what variables are most important for determining whether a credit card application will be approved or denied.

Error! Reference source not found.. Variable coefficient estimates, standard errors, and P-values.

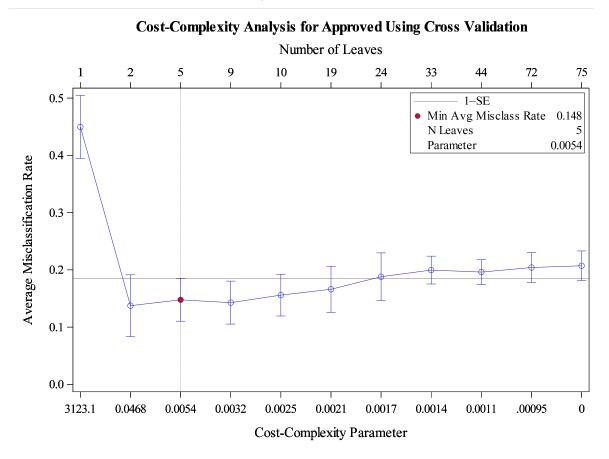
Analysis of Maximum Likelihood Estimates								
Parameter		DF	Estimate	Standard Error	Wald Chi-Square	Pr > ChiSq		
Intercept		1	2.7108	0.8244	10.8132	0.0010		
A4	I	1	17.8397	1689.7	0.0001	0.9916		
A4	u	1	0.8882	0.3224	7.5884	0.0059		
A4	у	0	0	•				
A6	?	1	-18.5497	1593.5	0.0001	0.9907		
A6	aa	1	-2.8463	0.8545	11.0952	0.0009		
A6	С	1	-2.3329	0.8115	8.2632	0.0040		
A6	СС	1	-1.2618	0.9473	1.7739	0.1829		
A6	d	1	-2.1218	1.0280	4.2604	0.0390		
A6	е	1	-1.9622	1.0636	3.4034	0.0651		
A6	ff	1	-4.2731	1.0310	17.1793	<.0001		
A6	\6 i		-3.2242	0.8971	12.9187	0.0003		
A6	j	1	-3.2649	1.3861	5.5480	0.0185		
A6	k	1	-2.9563	0.8881	11.0813	0.0009		
A6	m	1	-2.5419	0.9049	7.8906	0.0050		
A6	q	1	-2.1706	0.8490	6.5373	0.0106		
A6	r	1	-1.9058	3.8084	0.2504	0.6168		
A6	w	1	-1.8939	0.8474	4.9948	0.0254		
A6	x	0	0	-				
A9	f	1	-3.7630	0.3205	137.8444	<.0001		
A9	t	0	0	-				
A11		1	0.1644	0.0463	12.5904	0.0004		
A14		1	-0.00220	0.000881	6.2597	0.0124		
A15		1	0.000562	0.000182	9.5926	0.0020		

An alternative method that one might apply in this situation is a decision tree (Breiman et al. 1984). Decision trees (also known as classification and regression trees) work by recursive partitioning of the data into groups ("nodes") that are increasingly homogeneous with respect to some kind of a criterion, such as mean squared error for regression trees and either entropy or the Gini index

for classification trees. Ultimately the fitted tree is "pruned" back to remove branches and leaves of the tree that are just fitting noise in the data. The pruning process is a critical part of fitting a classification tree: unpruned trees *overfit* the data and are less accurate predictors for new data. The approach of segmenting the data space is quite different to that of fitting linear, quadratic or additive functions to the predictor variables. In cases where there are strong interactions among predictor variables, classification trees can outperform linear and quasi linear methods.

The first step in the fitting of a decision tree is to determine the appropriate *size* of the fitted tree. A plot of the cross-validated error rate against the size of the fitted tree is obtained using the code below:

Error! Reference source not found.. Cross-validated error plotted against the size (number of leaves) of the fitted trees.



The plot shows that the minimum cross-validated error rate is achieved by a tree with just 5 leaves, which is a very small tree, and the 1-SE rule of Breiman et al. (1984) selects a tree with just two leaves. That is, the tree splits the data just once. Usually, for large datasets, one would not expect such small trees to be effective predictors but for these data they are, and they provide us with some insight into the data.

The tree with just two leaves (terminal nodes branches) splits on the variable A9. Among the persons with a value of 't' on A9, 79.55% were approved for a credit card whereas of the persons with a value of 'f' on this variable, only 6.45% were approved for credit cards. One can only speculate as to what this question was, with 't' and 'f' being its only possible responses. The overall error rate for this simple split of the data is 13.74%, which is very comparable to the 12.6% for the logistic regression model.

How much can the error rate be reduced by using additional variables? The surprising answer is, "not much." The decision trees with 5 and 10 leaves have error rates of 14.36% and 14.44%, respectively, no better than—and , perhaps, a smidge worse than—the error rate for the simplest decision tree with just two leaves. Even random forests, one of the most accurate machine learning predictive methods, can only reduce the error rate to 12.5%. What this means is that nearly all of the information in these data about the approval or lack of approval of a credit card application is contained in the single variable, A9, and that a very simple decision tee identified this piece of information immediately.

## PREDICTION OF WINE QUALITY

The second example of applying machine learning methods for prediction concerns data on the quality of white wine in Portugal (Cortez et al. 2009). The response is the quality of the wine sample on a scale of 0—10, with 10 being the highest quality. The median value of the score of 3 experts was used. Predictor variables are chemical and physical characteristics of the wine samples including pH, density, alcohol content (as a percentage), chlorides, sulfates, total and free sulfur dioxide, citric acid, residual sugar, and volatile acidity.

In ordinary multiple linear regression we minimize the residual sum of squares,

$$RSS = \sum_{i=1}^{n} (Y_i - \beta_0 - \beta_1 x_{i1} - \beta_2 x_{i2} - \dots - \beta_p x_{ip})^2$$

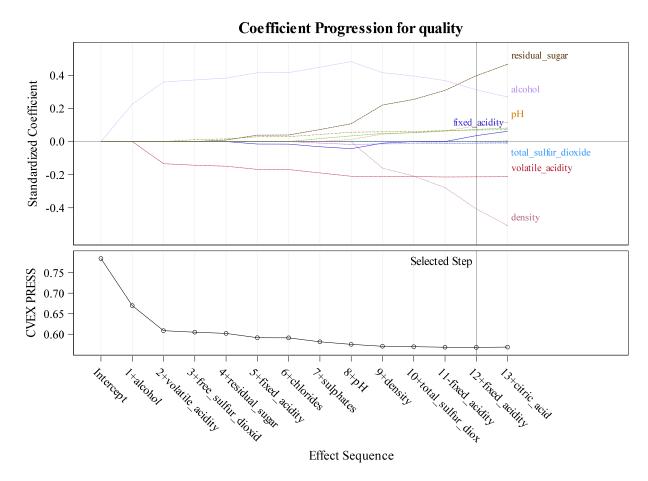
with respect to  $\beta_0, \beta_1, \beta_2, \cdots, \beta_p$  to obtain the least squares estimates of  $\beta_0, \beta_1, \beta_2, \cdots, \beta_p$ . The LASSO adds a penalty term to the residuals sum of squares. That is, we minimize

$$RSS = \sum_{i=1}^{n} (Y_i - \beta_0 - \beta_1 x_{i1} - \beta_2 x_{i2} - \dots - \beta_p x_{ip})^2 + \lambda \sum_{j=1}^{p} |\beta_j|.$$

The parameter  $\lambda$  is varied and a specific value may be chosen by some criterion, such as AIC or SBC, or to minimize cross-validated prediction error. In the SAS code below, the value of the LASSO parameter is selected by minimizing cross-validated error. Fifty distinct values of  $\lambda$  are tried. A plot of the coefficients as a function of  $\lambda$  follows the code.

run;

Error! Reference source not found.. Values of regression coefficients for different values of LASSO parameter  $\lambda$ .



From this plot we see that alcohol is the first variable to have a non-zero coefficient as  $\lambda$  decreases and volatile\_acidity is the second such variable. For this model the cross validated prediction error (CVEX PRESS) is 0.5679 and the final model contains all the predictor variables except citric acid. The LASSO estimates of the regression coefficients are given in Table 3. Increased

quality is associated with **larger** values of alcohol, residual sugar, and pH. Increased quality is associated with **smaller** values of density, chlorides, and volatile acidity. The coefficient for density is very large relative to the other regression coefficients but that is only a reflection of the fact that the differences in densities among the wines are very, very small. Fixing the random seed for the 10-fold cross-validation ensures that we are able to replicate results exactly when repeating the analysis.

Error! Reference source not found.. LASSO estimates of regression coefficients for white wine data.

Parameter Estimates						
Parameter	DF	Estimate				
Intercept	1	121.032634				
fixed_acidity	1	0.037306				
volatile_acidity	1	-1.873477				
residual_sugar	1	0.069450				
chlorides	1	-0.357294				
free_sulfur_dioxide	1	0.003608				
total_sulfur_dioxide	1	-0.000234				
density	1	-120.539600				
рН	1	0.553261				
sulphates	1	0.570229				
alcohol	1	0.224899				

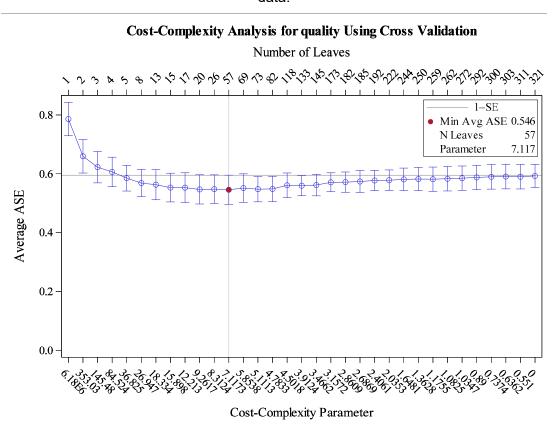
The second step in the analysis to fit a classification tree to the data. As was the case in the first example, the first step in fitting a classification tree is to determine how large the tree should be. Sample code for doing this is provided below:

By default PROC HPSPLIT "bins" the values of each numerical predictor variable into 100 bins of equal width across the range of the predictor variable. This is a small departure from the original algorithm of Breiman et al. (1984) in which the values of each numerical predictor variable are completely sorted. This modification makes perfect sense for very large datasets for which the cost of complete sorting would be prohibitive. For moderate sample sizes I prefer the original

algorithm and by selecting <u>intervalbins=10000</u> I am effectively making it so there are only 1 or 2 observations per bin, and hence coming close to a full sort of the predictor variables.

The plot of cross-validated error against tree size is given below. The cross-validated error is minimized for a large tree that has 57 leaves but the 1-SE rule of Breiman et al. (1984) selects a much smaller tree with only 5 leaves (Figure 5).

Error! Reference source not found.. Cross-validated error against tree size for the white wine data.



In figure 5 we see that at the root node, node 0, there are 4898 observations and the average quality score is 5.8779. The first split is on alcohol, at a value of 10.801. For the 3085 wines with alcohol < 10.801 the average quality score is 5.6055 whereas for the 1813 wines with alcohol  $\geq 10.801$  the average quality score is 6.3414. Thus the wines with higher alcohol content are rated higher, on average, and this result is consistent with the positive coefficient for alcohol in the regression. The difference between these two values may seem modest but the vast majority of the wines have scores in the range 5-8.

For the wines with alcohol < 10.801 the next split is on volatile\_acidity at a value of 0.250. The 1475 wines with volatile\_acidity < 0.250 have an average quality score of 5.8725 while the 1610 wines with volatile\_acidity  $\geq 0.250$  have an average score of 5.3609. This is consistent with the regression results in which volatile\_acidity had a negative coefficient.

The second split for the wines with alcohol  $\geq$  10.801, on free\_sulfur\_dioxide is much less interesting because only 114 out of the 1813 observations end up in the node corresponding to free\_sulfur\_dioxide < 11.012.

The cross-validated prediction error for the regression tree with 5 leaves is 0.5892, which is slightly larger than the value 0.5679 for the regression using LASSO estimates of the coefficients. By fitting a much larger regression tree, the prediction error may be reduced to 0.5485.

Error! Reference source not found.. First two levels of classification tree with 5 leaves (terminal nodes).

# **Subtree Starting at Node=0** Node Avg alcohol alcohol 1813 N Avg 5.6055 Avg 6.3414volatile\_acidity < 0.250 volatile\_acidity >= 0.250 free\_sulfur\_dioxide free sulfur dioxide < 11.012 $= 11.0\overline{1}2$ Node N Node Node Node 1475 Avg 5.8725 5.4123 Avg 5.3609 Avg Avg 6.4038

The third step in the analysis is to apply random forests to determine if higher predictive accuracy might be achieved. Random Forests (Breiman, 2001) takes predictions from many classification or regression trees and combines them to construct more accurate predictions. The basic algorithm is as follows:

- 1. Many random samples are drawn from the original dataset. Observations in the original dataset that are not in a particular random sample are said to be *out-of-bag* for that sample.
- 2. To each random sample a classification or regression tree is fit without any pruning.
- 3. The fitted tree is used to make predictions for all the observations that are out-of-bag for the sample the tree is fit to.
- 4. For a given observations, the predictions from the trees on all of the samples for which the observation was out-of-bag are combined. In regression this is accomplished by averaging the out-of-bag predictions; in classification it is achieved by "voting" the out-of-bag predictions, so the class that is predicted by the largest number of trees for which the observation is out-of-bag is the overall predicted value for that observation.

Many details are omitted from the discussion here, including the number of samples to be drawn from the original data, the size of those samples, whether the samples are drawn with or without replacement, and the number of variables available for the binary partitioning in each tree and at each node.

Random Forests may be fit using PROC HPFOREST in SAS® Enterprise Miner™. Here is some sample code for the white wine data:

All the predictor variables are interval valued and go in a single input statement. If there were categorical variables, we would need a second input statement for those variables with the option level=nominal. The response variable, Quality, is also interval-valued and goes into a target statement. The default number of subsets of the data and number of trees to fit is 200. The option scoreprole=oob asks for the out-of-bag error to be reported. One advantage of random forests over other machine learning algorithms is that nearly everything is automated and default settings produce good results in a large number of problems and settings. Table 4 below contains some accuracy results.

Error! Reference source not found.. Random Forests predictive accuracies for selected numbers of trees.

Fit Statistics							
Number of Trees	Number of Leaves	Average Square Error (Train)	Average Square Error (OOB)				
10	10472	0.08523	0.47121				
50	51571	0.06476	0.36657				
100	102746	0.06214	0.35143				
200	205276	0.06076	0.34452				

For the full 200 trees the out-of-bag average square error, which is equivalent to the cross-validated prediction error for multiple linear regression and regression trees, is 0.3445, which is quite a bit lower than the value of 0.5679 obtained from the regression using LASSO and the values of 0.5892 and 0.5485 obtained for the regression trees with 5 and 57 leaves, respectively. Thus, this is one situation where the use of a high level machine learning algorithm, such as random forests, gradient boosting machines, or support vector machines, can result in a much higher predictive accuracy than that which traditional regression methods could achieve.

A feature of random forests that is very popular with users is its algorithm for determining variable importance. Table 5 contains the variable importances for the white wine data. The data are sorted by the out-of-bag (OOB) mean squared error. The largest value is for alcohol at 0.05605 and there is a substantial drop to density at 0.02292. We decided to refit random forests with just the six most important variables. The out-of-bag error rate was 0.3709, which is a little higher than the value for the random forests fit with all 6 variables, but still much less than for the multiple linear regression model with LASSO estimation and the two regression trees.

Error! Reference source not found.. Variable importance from random forests analysis of white wine data.

Loss Reduction Variable Importance								
Variable	Number of Rules	MSE	OOB MSE	Absolute Error	OOB Absolute Error			
alcohol	24312	0.127562	0.05605	0.082900	0.031738			
density	19987	0.096659	0.02292	0.056390	0.005337			
volatile_acidity	13769	0.076166	0.01713	0.072654	0.033952			
free_sulfur_dioxide	19374	0.080078	0.00786	0.065752	0.016969			
chlorides	17315	0.066279	0.00099	0.043994	-0.000145			
citric_acid	15434	0.053505	-0.00788	0.053528	0.011781			
residual_sugar	16974	0.053933	-0.00982	0.051778	0.008112			
fixed_acidity	12515	0.044733	-0.01400	0.038773	0.001768			

Loss Reduction Variable Importance								
Variable	Number of Rules	MSE	OOB MSE	Absolute Error	OOB Absolute Error			
pH	20868	0.056580	-0.01843	0.058464	0.006669			
total_sulfur_dioxide	19861	0.059472	-0.01865	0.059930	0.007552			
sulphates	24667	0.055985	-0.02316	0.063720	0.007702			

What is the nature of the relationship between the individual predictors and the response variable? One approach to answering this is to construct partial dependence plots (Friedman 2000). An alternative is to fit a generalized additive model (Hastie and Tibshirani 1986). For an interval valued response we might assume an approximate normal distribution for the error terms and fit a GAM of the form:

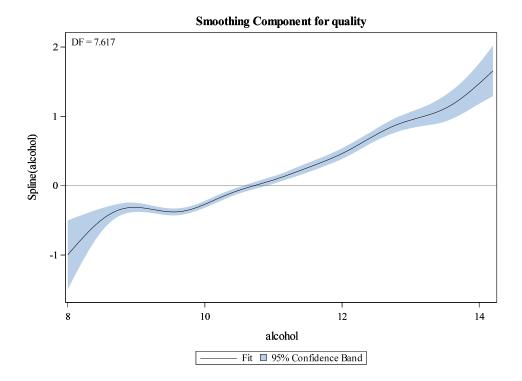
$$Y_i = s_0 + s_1(x_{i1}) + s_2(x_{i2}) + \dots + s_p(x_{ip}) + \varepsilon_i$$

where the  $s_1(\ ), s_2(\ ), \cdots, s_p(\ )$  are smooth functions of the respective predictor variables. This model is non-linear in that the  $s_j(\ )$  may be non-linear functions but it is also additive in the sense that no interactions among predictor variables are included. The primary output from fitting such models is a set of scatter plots of the  $s_j(x_{ij})$  against the values  $x_{ij}$ .

One procedure in SAS for fitting GAMs is PROC GAMPL. Sample code for fitting such a model to the white wine data using only the six most important variables identified by random forests follows:

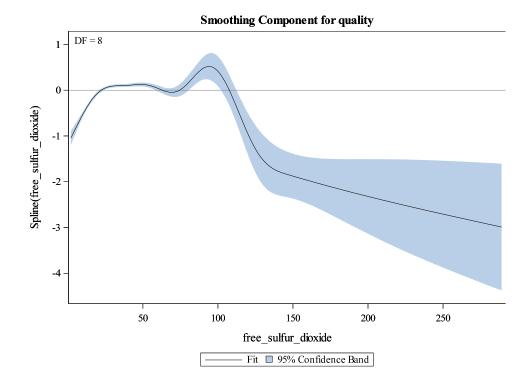
For the variable alcohol, identified by random forests as the most important variable for predicting quality, we obtain the following plot. This plot is remarkably linear.

Error! Reference source not found.. Transformation plot for alcohol in a GAM for the white wine data.



On the other hand, for free\_sulfur\_dioxide we obtain the following plot, which is noticeably non-linear.

Error! Reference source not found.. Transformation plot for free\_sulfur\_dioxide in a GAM for the white wine data.



For these data most of the transformation plots are approximately linear which raises the question: "If the relationships between most of the predictor variables and the response are (approximately linear), how is it that random forests significantly outperforms a multiple linear regression model for these data?" The answer lies in the fitting of interactions: the multiple linear regression model and even the generalized additive model do not incorporate interactions among predictor variables whereas tree-based methods including random forest and gradient boosting machines do.

# **CONCLUSION**

In this paper we have tried to illustrate the use of machine learning methods for regression problems. In the first example, on credit card applications, a simple decision tree captured the salient information that is available in the data. In the second example, random forests gave significantly more accurate predictions and generalized additive models provided a way to visualize the relationship between predictor variables and the response.

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